

Firms' Sales Expectations and Marginal Propensity to Invest*

Andrea Alati[†] Johannes J. Fischer[‡] Maren Froemel[§] Ozgen Ozturk[¶]

This version: 14th May 2024

Abstract

How do firms adjust their investment in response to sales shocks and what determines the response? Using a unique firm-level survey, we propose a novel approach to estimate UK firms' marginal propensity to invest (MPI) out of additional income: the forecast error of their sales growth expectations. Investment responds significantly to these sales surprises, with a 1 percentage point unexpected growth in sales translating into a 0.31 percentage point increase in capital expenditure. We find attentive firms to be more responsive, consistent with sales growth surprises providing firms with information about their demand. Sales growth surprises also cause firms to increase their prices, supporting this interpretation. We do not find evidence that these results are driven by financial frictions, uncertainty, or productivity shocks.

Keywords: Investment, survey data, corporate finance, financial frictions, learning

JEL codes: D22, D25, D84

*We thank Natalie Kessler, Simon Lloyd, Maarten De Ridder, participants at the 3rd workshop on Firm Dynamics, Market Structures and Productivity in the Macroeconomy, and the DMP team for helpful suggestions and comments. The views expressed here are those of the authors and not necessarily those of the Bank of England or its committees.

[†]Bank of England; e-mail: andrea.alati@bankofengland.co.uk

[‡]Bank of England; e-mail: johannes.fischer@bankofengland.co.uk

[§]Bank of England; e-mail: maren.froemel@bankofengland.co.uk

[¶]Department of Economics, University of Oxford; e-mail: Ozgen.Ozturk@economics.ox.ac.uk

1 Introduction

The dynamic of firms' investment is a crucial determinant of the resource allocation in an economy. It is therefore important to understand the drivers of firms' investment and firms' response to changes in their income in order to correctly quantify the effects of income shocks as well as those of macroeconomic policy on capital allocation.

A key measure to quantify these effects is the propensity of firms to invest out of unexpected income realisations, or their marginal propensity to invest (MPI). Identifying how firms' sales affect their gross capital investment (CAPEX) is not straightforward. On the one hand, firms might invest more today to generate more sales tomorrow. In that case, investment causes sales. On the other hand, firms might invest more today because they sold more yesterday and have more cash. In that case, sales cause investment.

In this paper, we provide an extensive investigation of firms' MPI and its drivers by using a novel approach that overcomes this identification issue. To identify unexpected changes in firms' income and estimate their MPI, we use a representative survey of UK firms' expectations, the *Decision Maker Panel*. This survey allows us to identify firm-level income surprises by constructing the forecast errors of firms' sales growth expectations, which we show are significantly correlated with firms' realised profitability. Importantly, these sales growth surprises are directly observable, fostering their construction without specific modeling assumptions.

We estimate the MPI as the change in investment induced by firms' own sales growth forecast errors. We find that a 1 percentage point higher than expected growth in sales translates into a 0.31 percentage point increase in investment. Consequently, there is a corresponding rise of 0.17 percentage points in the net asset stock. We also estimate the elasticity of firm-level price changes to sales growth surprises, where we find that an unexpected increase in sales growth by the same amount (1 percentage point) is associated with a 0.03 percentage point increase in prices over the subsequent 12 months.

The positive pass-through from sales surprises to investment and prices is consistent with a standard model of firm decision making under unobserved demand. When unexpected sales

realizations occur, they provide information about firms' actual demand. As a consequence, firms update their beliefs over time as they learn about their demand (as in [Berman, Rebeyrol and Vicard, 2019](#)). This learning process is reflected in their investment and price choices which, in this simple setting, can be shown to be linked to firms' expected sales.

We provide direct evidence for this learning-about-demand interpretation by showing that firms who are more attentive to their economic environment also respond more to sales growth surprises. This is because more attentive firms are able to extract more information from unexpected sales realisations. We measure attentiveness in two distinct ways. First, the DMP survey allows us to distinguish between firms that follow a *state-dependent* pricing model versus firms that instead follow a *time-dependent* pricing strategy. We estimate a significant MPI only for firms that follow state-dependent pricing. This is consistent with the assumption that state-dependent price-setters have to be more attentive to their economic environment. Better knowledge about their current state helps them to extract meaningful signals from unexpected sales realisation which translates to higher estimated MPIs. Second, we estimate a firm specific learning gain. Using these estimates, we find that firms with higher learning gains react more to sales growth surprises. Lastly, we show that firms facing higher levels of uncertainty ex ante are also more responsive, consistent with them drawing a larger information gain from the surprise.

To corroborate this interpretation we explore four alternative explanations for our results. First, we examine the potential role of financial frictions as a driving force behind the observed response. Second, we analyze the impact of (idiosyncratic) productivity shocks. Third, we assess whether firms exhibit heterogeneous reactions to micro and macro shocks as sources of sales growth surprises. Finally, we investigate uncertainty dynamics as a fundamental determinant of firms' forecast errors. We do not find support for these channels in the data which suggests that their role is small if anything.

We first explore the role of financial constraints, which can be an important factor shaping firm-level investment ([Khan and Thomas, 2013](#)) and increase firms' MPIs ([Jeenas, 2023](#)). Drawing on recent literature, we test for the role of financial constraints by estimating for heterogeneous MPIs along a number of commonly used proxies for financial constraints such

as age, size, leverage, liquidity, and interest burden (Cloyne, Ferreira, Froemel and Surico, 2023; Ottonello and Winberry, 2020; Jeenas, 2019). For none of these proxies we find any significant heterogeneity in the estimated MPIs. Taken together, these results suggest a limited scope for financial constraints to be driving our MPI measures.¹

An alternative explanation for our results could be that firm-level sales growth surprises are a proxy for productivity shocks, which change the optimal size of the firm and thus affect firm investment. We find that this is not the case: when we explicitly control for changes in firm productivity, our baseline estimate is remarkably stable. Furthermore, the estimated response of prices to sales growth surprises corroborates the interpretation. If firm-level productivity shocks had a more prevalent role in generating these surprises, standard monopolistic pricing would, in fact, imply a negative rather than positive relationship between productivity shocks and price changes. This suggests that our MPI measure captures a distinct aspect of firm behavior other than changes in productivity.

Furthermore, differential reactions of firms to micro and macro shocks, as in (Born et al., 2022), could affect our estimated MPIs. Through the decomposition of sales growth surprises into micro and macro components, we examine whether firms exhibit heterogeneous sensitivities to these shocks. Our findings indicate that firms do not adjust their gross investment differently in reaction to micro or macroeconomic news.

Finally, we investigate the role of firm-level uncertainty about future sales growth for our results. Firm-level uncertainty could explain our results in two different ways: On the one hand, elevated firm-level uncertainty about sales growth might be the result of large forecast error realisations and might lead to larger subsequent forecast errors. In this case, firms could be adjusting their investment in response to changes in uncertainty instead of unexpected sales realisation. On the other hand, the magnitude of firm-level uncertainty can shape how firms respond to sales shocks. This effect can go either way, with the real-options channel predicting a weaker response for high degrees of uncertainty, whereas firm-level learning predicting a stronger response. We find no evidence for the first effect because

¹Our results are consistent with Kaplan and Zingales (1997) who show how investment-cash flow sensitivities are a poor proxy of financial constraints, and with Ottonello and Winberry (2020) that find firms with low leverage responding more to monetary policy shocks than firms with high leverage.

our MPI estimates remain unchanged when controlling for current and past uncertainty. Furthermore, when interacting firm-level uncertainty with sales growth surprises, we find that relatively uncertain firms react significantly to sales growth surprises. Firms that are relatively certain, on the other hand, do not react. This is consistent with relatively more uncertain firms gaining more information from observing the current sales surprise and, thus, adjusting their investment more.

Taken together, our results suggest that the investment response to sales growth surprises is not driven by financial frictions, firm-specific productivity shocks, or differential response to micro and macro shocks, but stem rather from a behavioral response where income shocks help firms learn about the state of their demand, with more attentive firms extracting a strong and significant signal from unexpected income realisations.

Literature. Our paper is closely related to the empirical literature measuring the relationship between firms' capital investment and unexpected income shocks of various origin.

In a recent paper, [Martin-Baillon \(2021\)](#) applies the permanent/transitory shock decomposition developed in the consumption literature ([Blundell, Pistaferri and Preston, 2008](#)) to firms' income and estimates their investment response. In this case, the identification relies on correctly specifying firms' income processes to back out the transitory income shocks. In our case, as we have direct information on firms' expectations we can compute the unexpected component of sales growth realisation directly in the data without relying on a specific structure for firms' income processes.

Other contributions estimate the effects of income shocks due to macroeconomic policies. [Cummins, Hassett and Hubbard \(1994\)](#) study tax reforms, and more recently [Hebous and Zimmermann \(2021\)](#) estimate an investment elasticity to federal spending using variation in federal procurement contracts in the US. [Ottonello and Winberry \(2020\)](#) among many others analyze how investment responds to monetary policy.

Relative to this range of approaches, our main contribution is that we study how firms react to unexpected changes in income irrespective of the source of income variation. Recent research shows that firms react differently to, for example, micro and macro news ([Born](#)

et al., 2022). Our measure likely captures both types of news. Furthermore, we provide a comprehensive analysis of potential factors that could explain firms' MPIs. As part of this investigation, we propose a novel methodology to identify firm-level income shocks that does not rely on a model and yields more frequent income shocks than, e.g., tax reforms, which also have a larger magnitude than monetary policy shocks. The resulting MPI estimates are relatively large compared to the existing literature. We show that this is partially due to the fact that we can accurately control for firms' information set and expectations. This leads to larger estimates relative to the existing literature and highlights the importance of capturing firms' expectations.

Our paper is also closely connected to a few recent studies that use firms' sales expectations to understand firm behavior. [Barrero \(2022\)](#) establishes evidence on U.S. managers' belief biases to show that sales forecast errors are positively correlated with contemporaneous hiring decisions. In contrast with this work, we do not find evidence that firms over- or under-extrapolate. [Bachmann, Carstensen, Lautenbacher and Schneider \(2021\)](#) study the dynamics of sales uncertainty in a panel of German manufacturing firms, documenting a positive relationship between the magnitude of firms' sales growth forecast errors and their uncertainty about subsequent sales growth. Lastly, [Boutros et al. \(2020\)](#) report similar results based on a sample of forecasts by Chief Financial Officers of S&P 500 companies. Relative to these papers, our focus instead is on firms' investment decisions in response to unexpected sales growth. And, while we confirm the connection between sales forecast errors and subsequent sales growth in the latter two, this channel does not explain our main result. We also relate to the literature on measuring firm-level uncertainty and its interaction with firms' decision-making. A number of contributions have documented using aggregate and micro-level data that uncertainty can lower the level of economic activity. ([Bloom, 2009](#)). More specifically, higher uncertainty at the firm level can lower investment due to irreversibility, and weaken its responsiveness of investment to demand shocks (e.g. monetary policy) through a real options channel ([Bloom, Bond and Van Reenen, 2007](#); [Lakdawala and Moreland, 2022](#); [Bloom et al., 2022](#)). Recent work involving DMP data constructs subjective uncertainty measures to track economic uncertainty during the Covid-19 pandemic ([Altig](#)

et al., 2020) and finds that increased uncertainty contributed significantly to the negative impact of Brexit on UK firms (Bloom et al., 2019). Our results suggest that while uncertainty is not the main driver of our MPI estimate, firms with higher subjective uncertainty have a higher MPI, consistent with them having a larger informational gain from sales surprises.

Outline. The remainder of the paper is structured as follows: Section 2 describes the data. Section 3 describes our identification and presents stylised facts of sales growth surprises. Section 4 presents the empirical approach and the results of our empirical analysis. Section 5 discusses the main interpretation of our results and explores alternative explanations. Section 6 concludes.

2 Data

In this section, we provide an overview of the two primary data sources used in this paper for expectations (Decision Maker Panel) and balance sheet data (Bureau van Dijk). We then discuss the properties of our identified sales growth surprises.

2.1 The Decision Maker Panel

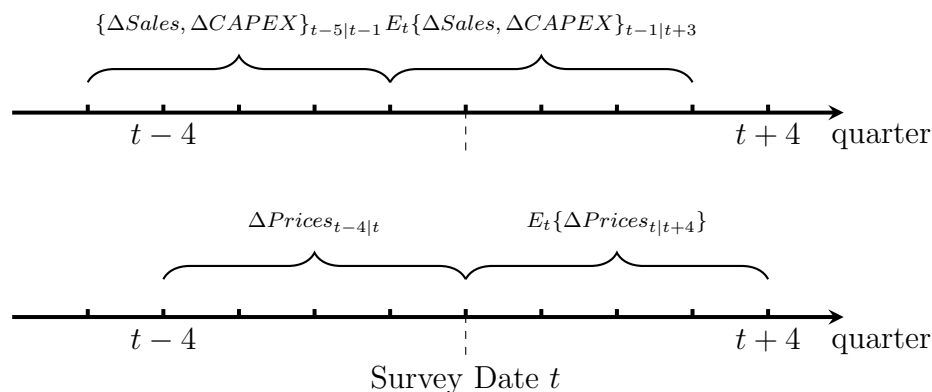
The DMP is a comprehensive and representative online survey focusing on UK businesses, jointly initiated by the Bank of England, the University of Nottingham, and Stanford University. With a monthly response rate approaching 3,000 firms, equivalent to about 5% of UK private sector employment, the DMP is one of the substantial monthly business surveys in Europe. The survey provides robust coverage across various industries, firm sizes, and regions in the UK for both public and private firms.

Since 2016Q3, the DMP surveys firms' expectations for one year-ahead gross investment as well as sales growth and own price growth.² In the first stage, firms are asked about the

²The survey employs a three-panel rotation system, wherein new participants are randomly assigned to one of three panels (A, B, or C). In any given month, each panel receives a third of the total questions, ensuring that over a quarter, all firms encounter the entire set of questions. This rotation method not only maintains a short survey for respondents but also ensures a consistent monthly influx of data.

expected realisation of each variable in the lowest, low, medium, high, and highest scenario (see Appendix A for the exact questions). Subsequently, firms are tasked with assigning probabilities to each of these scenarios. These probabilities allow us gauge firms’ confidence in these predictions. The expected value is then calculated using a weighted average of these scenarios, where the weights are the subjective probabilities assigned to each scenario. Additionally, the DMP surveys firms’ realised investment, sales and own price growth over the preceding year. Table B.1 provides summary statistics of firms’ survey responses.

Figure 1: Reference Periods for Survey Questions in Time t



As depicted in Figure 1, the DMP applies different reference periods for questions about CAPEX and sales on the one hand, and questions about prices on the other hand. For CAPEX, the period t survey asks about the *level* of firms’ investment in period $t - 1$. Firms are then asked about the realised (expected) level of CAPEX four quarters before (after) that reference quarter, i.e. $t - 5$ ($t + 3$). The growth rate of CAPEX is then computed using these respective levels.³ For prices, the period t survey asks about the *growth* of firms’ own prices using period t as reference period. That is, the survey asks firms about their realised (expected) own price growth between period t and four quarters before (after) that reference quarter, i.e. $t - 4$ ($t + 4$). Sales-related questions align with CAPEX in timing,

³In the DMP, the growth rate of CAPEX is computed using the same approach as in [Davis, Haltiwanger and Schuh \(1998\)](#), i.e. by normalising the change in CAPEX between $t - 1$ and $t + 3$ using the average CAPEX between $t - 1$ and $t + 3$. This weighted growth rate measure is preferable to standard growth rate measures (i.e. $\Delta \log(\text{CAPEX}_{i,t})$) because it does not get arbitrarily large when one of the elements approaches zero and does not produce missing observations when a firm does not invest in a given quarter. Furthermore, as [Baumeister and Hamilton \(2023\)](#) point out, this growth rate resembles a first-order Taylor approximation of $\Delta \log(\text{CAPEX}_{i,t})$ at the midpoint between the two elements. The approximation is almost exact as long as $\text{CAPEX}_{i,t}$ and $\text{CAPEX}_{i,t-1}$ do not differ by more than a factor of two.

using the previous quarter $t - 1$ as reference period. Unlike the CAPEX-related questions, sales-related questions focus on firms' sales growth instead of levels.

At a given survey date, price and sales as well as investment questions have a different reference period. This implies that the survey date for a given reference period also differs. For instance, firms' sales expectations for the period between t and $t + 4$ are surveyed in $t + 1$ whereas price expectations for the same period are surveyed in t . Expectations for prices thus reflect a different information set from that of sales and investment. This means that we cannot consistently deflate sales expectations with own price expectations and necessitates the use of nominal variables. Furthermore, measuring the unexpected income change in terms of quantities rather than pounds would also complicate the interpretation of the resulting estimate as a marginal propensity to invest out of unexpected income.

The survey drops all responses in which firms did not provide answers for all five scenarios or the order of scenarios is reversed. Furthermore, it drops observations in which the probabilities assigned to these five scenarios do not add up to 100%. Additionally, we drop observations in the utilities, finance and insurance, real estate, public, as well other services (such as Greenpeace) sectors as common in the literature.

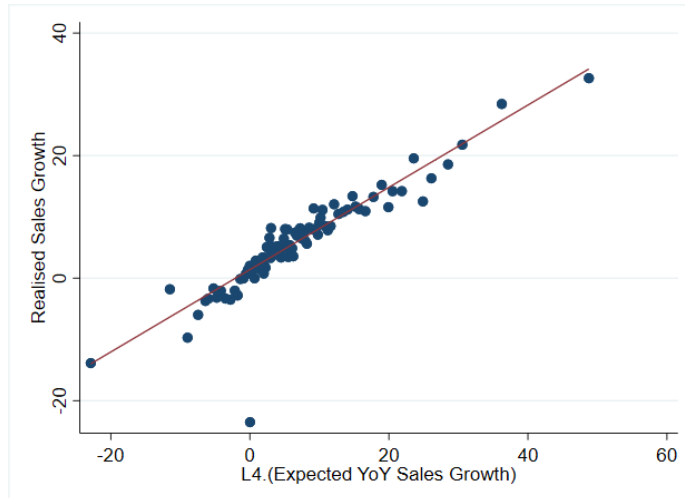
2.2 Balance Sheet Data

Firms in the UK are obligated to submit their annual accounts to Companies House, a government entity. These annual records are accessible through Bureau van Dijk (BvD) which compiles about 1.5 million distinct company accounts annually. Since BvD is a live database we use archived data sampled at a six-monthly frequency, ensuring that we gather information in the initial form it appeared in BvD database (see [Bahaj, Foulis and Pinter, 2020](#)). Importantly, balance sheet data provided by the BvD can be merged with DMP data through a shared firm identifier. This allows us to compute relevant balance sheet information (i.e., leverage, net liquidity, interest burdens, profitability), as well as firms' age and productivity growth. Table B.2 displays the summary statistics of these measures.

3 Sales Growth Surprises

As illustrated in Figure 2, firms’ sales growth expectations reported in the DMP are highly correlated with their respective realisations ($\rho = 0.72$). Despite this high correlation, firms’ sales expectations can explain relatively little variation of sales realisations ($R^2 = 0.13$), indicating that firms frequently make large forecast errors.

Figure 2: Sales Expectations and Realisations



Note: This figure plots the bins of sales expectations & subsequent sales realisations along with a linear regression line which has a slope of $\rho = 0.72$ ($\sigma = 0.03$, $R^2 = 0.13$).

We use these forecast errors to identify shocks to firms’ incomes. For completeness, define sales growth surprises as the difference between a) firm i ’s realised sales growth between $t - 4$ and t as reported in period $t + 1$ and b) firm i ’s sales growth expectations for the same 4-quarter window as reported in period $t - 3$:

$$FE_{i,t-3}(\Delta\text{Sales}_{i,t-4|t}) = \Delta\text{Sales}_{i,t-4|t} - E_{i,t-3}\Delta\text{Sales}_{i,t-4|t} \tag{1}$$

These identified sales growth shocks combine three attractive features. First, they are directly observable in the data and do not require assumptions about the firms’ income process. Second, these shocks are both frequent and sizable, setting them apart from other aggregate and firm-level shocks used in the literature. Third, the exposure to sales growth surprises is

equal across firms, unlike the exposure to aggregate shocks like monetary policy. In rest of this section, we discuss properties of these sales growth surprises and present stylised facts.

Table 1: Summary Statistics

	N	Mean	s.d.	p10	p50	p90
$\mathbb{E}_{t+1}\Delta\text{Sales}_{t t+4}$	53,983	7.32	10.76	-3.00	5.00	21.45
$\sigma(\Delta\text{Sales}_{t t+4})$	53,983	5.90	4.49	1.49	4.66	12.77
$\Delta\text{Sales}_{t-4 t}$	55,993	6.62	20.89	-19.00	5.00	34.00
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	21,768	-0.72	20.39	-24.79	-0.20	23.00

1. *Firms' expectations are unbiased but overprecise.* Table 1 shows that forecast errors are zero on average. That is, firms are not systematically under- or overoptimistic.⁴ However, firms underestimate the volatility of sales growth, i.e. their sales expectations are overprecise. This is illustrated by the difference between subjective sales growth uncertainty $\mathbb{E}_{t+1}(\sigma(\Delta\text{Sales}_{t|t+4}))$ and the volatility of realised sales growth $\Delta\text{Sales}_{t-4|t}$ in Table 1. This difference is further illustrated in Panel A of Figure 3, which plots the histogram of subjective vs. realised forecast errors. Both observations are consistent with [Bachmann et al. \(2021\)](#) and [Barrero \(2022\)](#) who use the forecast errors of firms' sales expectations in Germany and the U.S., respectively.

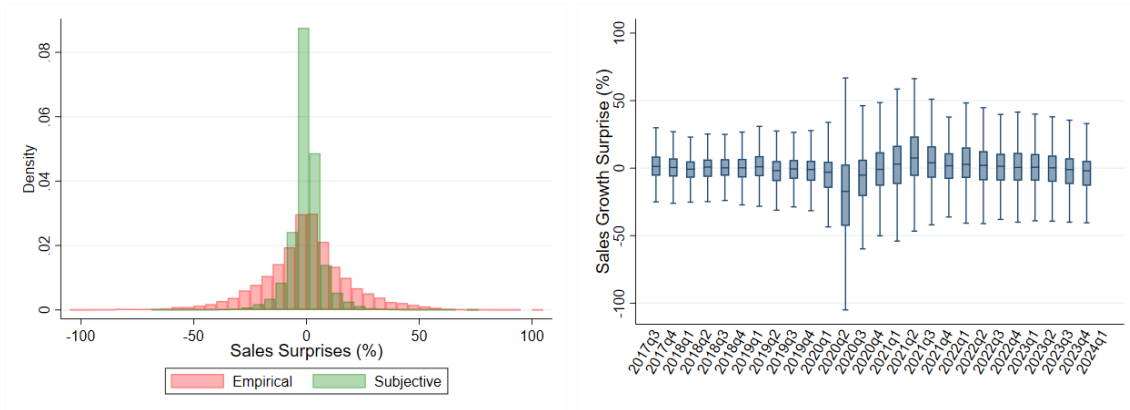
2. *Firms do not make persistent forecast errors.* Firms' forecast errors for year-on-year sales growth are not correlated with the forecast error for the same window in the previous year, i.e. forecast errors are not correlated within firm over time as Column (1) of Table B.3 shows. Presumably this is because firms do not under- or overextrapolate from past sales realisations so that sales forecast errors are not significantly related to past sales growth (expectations) as Columns (2)-(3) show. In this, our findings differ from [Barrero \(2022\)](#).

3. *Firms' forecasts get better over time.* Column (4) of Table B.3 shows that the magnitude of forecast errors decays over time, indicating that firms improve their forecasts.

4. *Firms under-react to news.* Table B.4 shows that firms' forecast revisions are positively related to subsequent forecast errors, indicating that firms underreact to news. Decomposing this forecast revision into a industry-specific and an idiosyncratic component by regressing

⁴This result still holds when correcting the standard error of the average forecast error by clustering on the firm as well as on the industry and quarter level.

Figure 3: Distribution of Sales Growth Surprises



(a) Histogram of Sales Growth Surprises

(b) Cyclicity of Sales Surprises

Note: Panel a of this figure plots the histogram of realised sales growth surprises (red) along with a hypothetical distribution of sales growth forecast errors (green) that follows firms’ subjective probability distributions (i.e. a histogram of sales growth surprises obtained by taking independent draws from firms’ subjective probability distributions). Panel b of this figure includes the box plot depicting the median sales growth surprises in each quarter along with the interquartile range as well as the minimum and maximum points of the distribution that omits outliers (values that are $1.5 \times \text{IQR}$ above the upper quartile/below the lower quartile).

it on quarter \times industry dummies furthermore shows that firms primarily under-react to idiosyncratic news. The estimated coefficient on industry-specific news is negative but insignificant, which suggests that firms adjust their expectations correctly to these macro news. In this, our findings depart from [Born et al. \(2022\)](#) who find evidence for an under-reaction of own price growth expectations to macro news by German firms.

5. Forecast errors are cyclical but largely idiosyncratic. Panel B of Figure 3 illustrates the cyclical behaviour of forecast errors: with the onset of the Covid-19 pandemic, the share of firms experiencing positive forecast errors drops sharply before recovering in 2021. Similarly, the dispersion of forecast errors spikes around the onset of the Covid-19 pandemic and remains elevated compared to pre-pandemic periods. However, even in 2020Q2 a substantial share of firms reported higher sales growth than they expected before the pandemic and in all periods the interquartile range of sales growth surprises includes zero. This indicates the largely idiosyncratic nature of sales growth surprises.

6. Sales growth surprises are associated with changes in contemporaneous profits. The DMP does not survey firms’ profit expectations so that we cannot compute profit surprises.

Nonetheless, Table 2 shows that sales growth surprises are significantly correlated with firms' profits & profitability: a 1pp sales growth surprise is associated with a contemporaneous 0.51pp growth in profits (EBITDA) and a 0.04 higher profitability (EBITDA-to-assets).⁵

Table 2: Sales Surprises and Balance Sheet Data

	(1)	(2)
	$\Delta\text{EBITDA}_{t-4 t}$ (BvD)	$\Delta\text{Profitability}_{t-4 t}$ (BvD)
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.51*** (0.14)	0.04** (0.02)
$\Delta\text{Price}_{t-4 t}$	1.49** (0.75)	0.24*** (0.08)
N	711	796

Note: This table reports the contemporaneous relationship between the growth rate of EBITDA & EBITDA-to-Assets (reported in BvD) and sales growth surprises. All estimates in Table are based on the same sample of firms as in Table 4, Column (1), but rely only on periods in which firms filed their balance sheet information. Standard errors in Columns (2) and (3) clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

7. *Forecast errors predict a non-linear change in subsequent sales growth (expectations).* Column (1) of Table 3 shows that firms, on average, revise their sales expectations downwards in response to a positive sales growth surprise.⁶ However, this average response masks significant heterogeneity in responses to negative and positive sales growth surprises, as Column (2) shows (see Figure B.1 for a visualization of this non-linearity). The response to negative surprises is also more persistent, as Table B.5 shows. Furthermore, columns (3) and (4) of Table 3 indicate that firms' adjustment of sales expectations aligns very well the eventual change in sales growth.

⁵All estimates in Table 2 are based on the same sample of firms as in Table 4, Column (1), but relying only on periods in which firms filed their balance sheet information.

⁶We control for firms' price growth expectations plans for a potential price response in a regression of sales growth expectations on sales forecast errors.

Table 3: Sales Surprises and Future Sales

	(1)	(2)	(3)	(4)
	$\mathbb{E}_{t+1}\Delta\text{Sales}_{t t+4}$	$\mathbb{E}_{t+1}\Delta\text{Sales}_{t t+4}$	$\Delta\text{Sales}_{t t+4}$	$\Delta\text{Sales}_{t t+4}$
	b/se	b/se	b/se	b/se
$FE_{t-3}(\Delta\text{Sales}_{t-4 t})$	-0.10*** (0.02)		-0.08** (0.03)	
$FE_{t-3}(\Delta\text{Sales}_{t-4 t})^-$		-0.44*** (0.06)		-0.38*** (0.07)
$FE_{t-3}(\Delta\text{Sales}_{t-4 t})^+$		0.06** (0.03)		0.02 (0.05)
$\mathbb{E}_t(\Delta\text{Price}_{t t+4})$	0.46*** (0.09)	0.43*** (0.09)	0.44*** (0.15)	0.41*** (0.15)
Negative = Positive (p-value)		0.00		0.00
N	4,263	4,263	3,701	3,701

Note: This table reports the relationship between sales expectations (Columns (1) and (2)) and realisations (Columns (3) and (4)) and the preceding sales forecast errors (split between positive and negative realisations in Columns (2) and (4)) in the sample used to estimate model (2). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

4 Empirical Results

Having documented the informational content of our identified sales growth surprises for contemporaneous profits and subsequent sales growth, we now estimate how firms respond to these sales growth surprises. We focus on estimating the marginal propensity to invest in Section 4.1. We provide additional evidence on the response of prices in Section 4.2.

4.1 The Average MPI

We estimate the MPI as the realised change of firm i 's gross investment between t and $t + 4$ as a function of the unexpected realisation of sales growth in previous year, i.e. the forecast error between $t - 4$ and t . That is, we estimate the MPI as the coefficient $\hat{\beta}$ using

$$\begin{aligned} \Delta\text{CAPEX}_{i,t|t+4} = & \alpha + \beta FE_{i,t-3}(\Delta\text{Sales}_{i,t-4|t}) + \gamma_1 \Delta\text{Prices}_{i,t-4|t} \\ & + \gamma_2 \Delta\text{CAPEX}_{i,t-4|t} + \gamma_3 E_{i,t+1}(\Delta\text{Sales}_{i,t|t+4}) + u_{i,t}, \end{aligned} \quad (2)$$

The set of controls includes realised price growth to control for the endogenous part of the forecast error. This eliminates the component of the sales surprise that is caused by firms adjusting their prices during the forecast horizon. For example, if firms experience higher sales growth than expected in period $t - 3$ because they increased their prices in $t - 2$ by more than originally planned, the forecast error realisation would no longer be unexpected by the firm. Furthermore, controlling for realised price growth partials out any supply driven source of the forecast error, allowing us to interpret $\hat{\beta}$ as the response to demand driven sales surprises. To the extent that firms contemporaneously observe sales growth different than expected in period $t - 3$ and adjust investment or prices before period t , the estimated MPI $\hat{\beta}$ would constitute a lower bound of the response to demand induced sales shocks. We further control for past investment growth to account for lumpiness in investment. We do not include firm fixed effects on top of lagged investment to avoid biasing our result (see [Nickell, 1981](#)). Finally, we control for expected future sales growth to control for potential heterogeneity in the persistence of sales growth surprises across firms.⁷ This partials out the immediate effect of sales growth surprises on firms' expectations.

To assess the robustness of our MPI measure we estimate three alternative versions of model (2): First, given that our main sample includes the COVID years, a potential concern is that our estimates of MPIs are driven by the unprecedented nature of the COVID shock. To address this concern, we estimate (2) on a restricted sample in which we drop the sales realisations during the initial COVID quarters (2020Q2 and 2020Q3), excluding the variation in our main explanatory variable induced by the pandemic. Second, we also augment the specification in Equation (2) with industry \times quarter fixed effects to control for aggregate shocks.⁸ Finally, to avoid the possibility that firms' forecasts of tail scenarios might be biasing our MPI estimate, we run the estimation using an alternative measure of sales surprises based on firms' median forecasts rather than their probability weighted ones.

⁷In Table C.10 we show that the shocks identified via sales growth surprises do not exhibit any heterogeneous persistence across firms even when conditioning on firms' attentiveness types.

⁸Here, industry refers to 12 high level sectors: agriculture & forestry & fishing, mining & quarrying, manufacturing, construction, wholesale & retail, transport & storage, accommodation & food, information & Communication, professional & scientific & technical activities, administrative & support service activities, health, as well as arts & entertainment & recreation.

Table 4: Average MPI

	(1)	(2)	(3)	(4)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.31*** (0.08)	0.41*** (0.09)	0.28*** (0.09)	
$\text{FE}_{t-3}(\text{med}(\Delta\text{Sales}_{t-4 t}))$				0.30*** (0.08)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.01*** (0.11)	0.90*** (0.15)	0.70*** (0.12)	
$\mathbb{E}_{t+1}(\text{med}(\Delta\text{Sales}_{t t+4}))$				1.12*** (0.12)
$\Delta\text{CAPEX}_{t-4 t}$	-0.25*** (0.02)	-0.24*** (0.02)	-0.25*** (0.02)	-0.25*** (0.02)
$\Delta\text{Price}_{t-4 t}$	-0.13 (0.38)	0.03 (0.38)	-0.63* (0.38)	-0.15 (0.38)
Covid	Yes	No	Yes	Yes
FE	No	No	Quarter x Ind.	No
N	4,312	3,918	4,290	4,312

Note: Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. $\mathbb{E}_{t+1}(\text{med}(\Delta\text{Sales}_{t|t+4}))$ refers to the forecast of sales growth in the median scenario whereas $\text{FE}_{t-3}(\text{med}(\Delta\text{Sales}_{t-4|t}))$ refers to the forecast error for sales growth in the median scenario.

Results. We report the main estimates for the average MPI in Table 4. Column (1) shows our preferred measure estimated on the full sample of firms, including the COVID quarters.⁹ Our estimate indicates that, on average, a 1 percentage point unexpected growth in sales translates to a 0.31 percentage point increase in gross investment. More concretely, if a (hypothetical) firm expected to sell £100 worth of goods and planned to invest £10 happens to actually sell £110 (a 10% surprise), it will increase investment by £3.10 from £10 to £13.10 over the following year.

The estimated MPI is remarkably robust across different specifications, as shown in Columns (2) to (4) of Table 4. Excluding the COVID quarters increases our MPI estimate to 0.41 (Column (2)), while including quarter \times industry fixed effects to control for aggregate shocks marginally reduces the point estimate to 0.27 (Column (3)). Reassuringly, Column (4) shows

⁹This setup requires survey responses in time $t - 3$ (sales growth expectations), t (realised price growth), $t + 1$ (realised sales and investment growth), and $t + 5$ (realised investment growth). Taking into account these leads and lags, the effective sample of sales growth surprises spans 2017q3 - 2023q2, with data on realised investment covering the period 2018q3 - 2024q2.

that measuring sales surprises using firms' median forecasts rather than their probability weighted ones delivers exactly the same estimate as in the baseline specification. None of these alternative estimates deviates significantly from our baseline specification.

The remaining coefficient estimates have the expected sign and are similarly stable across specifications. Gross investment is significantly positively correlated with sales growth expectations, indicating firms increase their productive capacity when expecting higher sales. Gross investment is negatively correlated with its lagged realisation, reflecting the lumpy nature of capital investments at the firm level.

Comparison with the literature The two papers closest to ours are [Martin-Baillon \(2021\)](#) and [Hebous and Zimmermann \(2021\)](#), who estimate the investment response of public U.S. firms to transitory sales shocks to lie between 10 – 15%, between a third and a half of our estimate. A significant part of this difference comes from the fact that we can accurately capture firms' information sets and expectations. That is, transitory sales shocks change firms' sales growth expectations. We estimate this effect to be negative (see Table 3). This negative association implies that positive sales growth surprises depress firms' expectations of future sales growth, subsequently dampening investment. Therefore, not controlling for sales growth expectations leads to a smaller estimated investment elasticity because the estimated elasticity is a composite of the direct response and expectation adjustment. In our case, the estimated elasticity drops from 0.31 to 0.21 if we omit the expected sales growth control (see Table C.1). This underscores the importance of precisely capturing firms' expectations. The remaining difference is likely attributed to variations in shock identification methods and sample set, particularly the inclusion of privately owned firms.

Sensitivity. To corroborate our results, we re-estimate our baseline specification (2) but replace gross investment realisations surveyed in the DMP as dependent variable with balance sheet data on realised net investment from annual balance sheet accounts. In particular, in Table C.2 we consider three different measures of investment: the growth rate of tangible assets, the growth rate of fixed assets, and the growth rate of intangible assets. The results in Columns (1) and (2) indicate that sales growth surprises not only lead to an increase

in gross investment, but also lead to a significant increase in the net fixed (tangible) asset stock. Finally, Column (3) indicates that sales growth surprises do not lead to changes in the intangible asset stock. Taken together, these results indicate that the estimated MPI reflects neither replacement investment due to utilization-dependent depreciation rates nor inventory building. The latter would also be at odds with the phrasing of the survey question.

Table C.3 shows the effect of sales growth surprises on investment growth over a medium-term horizon: The response of investment growth between $t + 4$ and $t + 8$ to a sales growth surprise between $t - 4$ and t is negative and marginally significant. This indicates that investment growth shows some tendency to mean revert but the level of investment appears to remain elevated also in the medium term.¹⁰

Note again that these results are obtained from nominal variables due to the difference in information sets underpinning firms' sales and price expectations. However, the results are basically unchanged even when we ignore the difference in information sets and use real variables instead (see Table C.4). To deflate firms' sales growth expectations (realisations) we use their own price growth expectations (realisations) and we deflate realised investment using the CPI. Furthermore, this average estimated MPI appears to be largely driven by large, positive sales growth surprises, whereas medium sized or large negative surprises do not lead to significant investment responses (see Table C.5). Table C.6 shows that the estimated MPI does not differ significantly between firms in the manufacturing & construction industries and firms in the services industries.¹¹

These results are also also robust to controlling for i) how long a given firm has been taking part in the survey, ii) whether the survey questions are answered by a given firm's CEO, CFO, or other personnel,¹² and iii) the legal form of the company (see Table C.7). Finally, we test whether our results change when weighting observations using employment weights and when using the sales growth surprise as an instrumental variables for realised sales growth

¹⁰A back of the envelope calculation suggests that investment is $(1 + 0.0032)(1 - 0.0017) = 0.15\%$ higher in $t + 8$ than in t .

¹¹We count the following industries as the service sector: wholesale & retail, transport & storage, accommodation & food, information & communication, professional & scientific & technical activities, administrative & support service activities, health, as well as arts & entertainment & recreation.

¹²For this control, we replace missing information on the survey respondent's position with the first / last available information.

instead of adding it directly to the estimated model. Table C.8 shows that in neither case the results change significantly.

4.2 The Price Response (MPMP)

In addition to the investment response, we also investigate how firms adjust their prices in response to unexpected changes in sales growth. For this, we use the same specification as in Section 4.1 but instead of investment growth we estimate the response of realised price growth between t and $t + 4$ to a sales growth surprise. The estimated coefficient measures how much the unexpected change in sales growth induces a change in future prices, providing an estimate of firms’ “marginal propensity to modify prices” (MPMP).

Table 5: Average MPMP

	(1)	(2)	(3)	(4)
	$\Delta\text{Price}_{t t+4}$	$\Delta\text{Price}_{t t+4}$	$\Delta\text{Price}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.03*** (0.01)	0.03*** (0.01)	0.02*** (0.00)	
$\text{FE}_{t-3}(\text{med}(\Delta\text{Sales}_{t-4 t}))$				0.03*** (0.01)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	0.07*** (0.01)	0.09*** (0.01)	0.05*** (0.01)	
$\mathbb{E}_{t+1}(\text{med}(\Delta\text{Sales}_{t t+4}))$				0.07*** (0.01)
$\Delta\text{CAPEX}_{t-4 t}$	0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	0.00 (0.00)
$\Delta\text{Price}_{t-4 t}$	0.45*** (0.03)	0.44*** (0.03)	0.33*** (0.03)	0.45*** (0.03)
Covid	Yes	No	Yes	Yes
FE	No	No	Quarter x Ind.	No
N	3,704	3,381	3,685	3,704

Note: Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. $\mathbb{E}_{t+1}(\text{med}(\Delta\text{Sales}_{t|t+4}))$ refers to the forecast of sales growth in the median scenario whereas $\text{FE}_{t-3}(\text{med}(\Delta\text{Sales}_{t-4|t}))$ refers to the forecast error for sales growth in the median scenario.

We report our baseline estimate in Table 5, Column (1). As for our MPI estimate, we control for the role of COVID, the possible influence of industry specific time trends and the regularity of firms’ expectations by changing our estimation period, Column (2), including

industry-quarter fixed effects, Column(3), and by using firms' median scenarios to construct the sales growth surprises, Column (4). As for the baseline MPI estimates, our measurement of firms' MPMP is very stable across the different specifications, indicating that a 1 percentage point higher than expected sales growth induces a 0.03 percentage point increase in prices, in the year following the realisation of the sales growth surprises. Note that the stable coefficient in column (3) despite the inclusion of quarter \times sector fixed effects suggests that the price response is not driven by the pass-through of higher prices of suppliers that are subject to correlated sales surprises.

5 Drivers

The average response to sales growth surprises can reflect a myriad of competing factors. In this section we discuss the main driver of our MPI measure: firms use unexpected sales growth realisations to learn about the slope of their demand curve. We also explore other dimensions of firm heterogeneity to rule out alternative explanations for the positive relationship between sales growth surprises and investment.

5.1 Attention

If firms use unexpected sales growth realisations to learn about the slope of their demand curve, we would expect more attentive firms to adjust their investment and prices by more than less attentive firms. The degree of attention to the state of the economy is of course an unobservable firm characteristic. However, the DMP provides us with two measures that can be used to proxy firm attentiveness: first, we utilise the DMP survey responses to estimate a firm specific learning gain. Second, the DMP directly asks whether firms follow a *state-* or *time-* dependent pricing strategy. Using these two measures, we show in this section that more attentive firms indeed respond more sales growth surprises. Importantly, this is consistent with our main result. Although we control for sales expectations in our baseline specification (2), which should capture the adjustment firms make following a sales surprise,

we find that firms under-react to sales growth surprises and their expectations adjust only sluggishly (see stylised fact 4 and Table B.4).

5.1.1 Learning Gain

We use the degree to which firms adjust their expectations in response to forecast errors, i.e. their learning gain (Evans and Honkapohja, 2001), as a proxy for attentiveness. We estimate the firm specific learning gain as the parameter γ_i that links quarter-on-quarter revisions in sales growth expectations to forecast errors from the following firm-level regression

$$E_{t+1}(\Delta\text{Sales}_{i,t|t+4}) - E_t(\Delta\text{Sales}_{i,t-1|t+3}) = \alpha_i + \gamma_i \text{FE}_{i,t-3}(\Delta\text{Sales}_{i,t-4|t}) \quad (3)$$

$$+ \zeta_i \Delta\text{Prices}_{i,t-4|t} + \delta_i \Delta E \Delta\text{Prices}_{i,t-4|t} \quad (4)$$

We only keep estimates of $\gamma_i \in (-1, 1)$ that were estimated on a sample of 8 or more observations. With this regression setup we deviate from the learning literature in two ways. First, we control for realised price growth as in our baseline model to account for the endogenous part of the forecast error. Second, we control for the change of price growth expectations because, upon realisation of the forecast error, firms jointly adjust their price and sales growth expectations. Controlling for price growth expectations accounts for any heterogeneity in this adjustment. Figure 4 plots the distribution of estimated learning gains. The average and median gain is negative (consistent with the negative relationship between sales growth surprises and subsequent sales growth documented in Table 3) with large outliers.¹³

To investigate whether firms' with different degrees of attentiveness respond differently to sales growth surprises, we create a dummy variable $D_{i,\tau}$ indicating whether $\gamma_i \in (-1, -1/5)$, $\gamma_i \in (-1/5, 1/5)$, or $\gamma_i \in (1/5, 1)$.¹⁴ Table C.9 reports summary statistics of firms by the magnitude of their respective learning gain. We add this interaction term to Equation 2 to

¹³The size of the learning gain is not related significantly with the magnitude of the forecast errors.

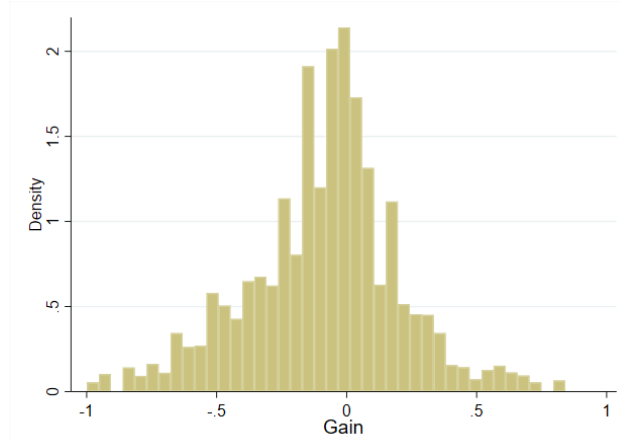
¹⁴Figure C.1 shows that the distribution of realised sales growth surprises is very similar for the three groups: All three distributions are centered around zero, but the distribution of sales growth surprises for firms with larger learning gains has more mass around the center.

estimate different MPIs for firms with different learning gains. In particular, we estimate:

$$\begin{aligned} \Delta\text{CAPEX}_{i,t|t+4} = & \alpha + \sum_{\tau} (\beta_{\tau} F E_{i,t-3}(\Delta\text{Sales}_{i,t-4|t}) \times D_{i,\tau} + \zeta_{\tau} D_{i,\tau}) + \gamma_1 \Delta\text{Prices}_{i,t-4|t} \\ & + \gamma_2 \Delta\text{CAPEX}_{i,t-4|t} + \gamma_3 E_{i,t+1}(\Delta\text{Sales}_{i,t|t+4}) + u_{i,t}, \end{aligned} \quad (5)$$

The coefficients β_{τ} estimate how the MPI varies across the firm distribution. This specification, has the advantage that it does not impose linearity on the interaction. Instead, it offers a non-parametric way of estimating the heterogeneous responses to sales growth surprises by different firm characteristics.

Figure 4: Histogram of Learning Gains



Note: This figure displays the histogram of firm-level learning gains estimated from model (4), restricting the estimated gains to lie in $\gamma_i \in (-1, 1)$.

Table 6 presents the results. We find that firms with large, positive learning gains, i.e. those that revise their expectations more in response to forecast errors, adjust their investment by more in response to sales growth surprises than firms with a small or large, negative learning gain. This difference is statistically significant. For completeness, Column (2) reports the estimated price response to sales growth surprises for the three types of firms. Here, we do not find any statistically significant heterogeneity for different levels of firm attentiveness.¹⁵

¹⁵These results remain unchanged when also interacting sales growth surprises with the size of the learning gain to account for a potentially heterogeneous updating of expectations between the three groups.

Table 6: Does Attentiveness Matter?

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
$\gamma_i \in (-1, -1/5) \times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.19 (0.20)	0.05*** (0.01)
$\gamma_i \in (-1/5, 1/5) \times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.20 (0.15)	0.03*** (0.01)
$\gamma_i \in (1/5, 1) \times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.82** (0.33)	0.01 (0.02)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.01*** (0.17)	0.08*** (0.01)
$\Delta\text{CAPEX}_{t-4 t}$	-0.27*** (0.02)	-0.00 (0.00)
$\Delta\text{Price}_{t-4 t}$	-0.05 (0.51)	0.44*** (0.03)
Low = Medium (p-value)	0.95	0.10
Medium = High (p-value)	0.08	0.39
Low = High (p-value)	0.09	0.08
N	2,763	2,593

Note: This table reports the results from estimating model (5), controlling for dummies for the respective interaction terms (omitted for brevity). The firm-level learning gain is estimated using model (4). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. See Tables B.6 & B.7 for the respective median value of the firm characteristic and median sales growth surprises for each group.

5.1.2 Price setting type

As an alternative measure of attention we estimate whether firms' differ along their price-setting strategy. The DMP reports whether firms follow a *state-* or *time-* dependent pricing strategy.¹⁶ Firms are asked whether they reset their prices in regular intervals, i.e. a time-dependent strategy, or in response to changes in economic conditions, i.e. a state-dependent strategy. Firms that follow a state-dependent pricing strategy have pay more attention to the state of their activities. As a consequence, they should react more to sales growth surprises

¹⁶This categorization is self reported in the DMP and was surveyed between 2023m2 - 2023m4 as well as 2024m2 - 2024m3. We assume that that the price-setting strategy it time-invariant. If, following a rational (in-)attention logic, some firms switched from a time-dependent to a state-dependent price setting strategy during the high inflation period of 2022-2024, this would imply that we underestimate the difference between the two types of firms because we wrongly classify firms as state-dependent price setters, i.e. attentive, even during the low inflation part of the sample.

as for these firms unexpected realisations of the state should contain a larger signal about the state of the economy relative to time-dependent pricing firms. Table C.11 reports summary statistics of firms by their respective price-setting strategy.

Table 7 reports the estimates of MPIs for firms that follow either a state-dependent or a time-dependent pricing strategy. As shown in Column (1), the response of CAPEX for sales growth surprises is strongly positive and significant only for firms that are state-dependent and therefore are structurally more versed in responding to changes in their economic environment. For these firms, a 1 percentage point higher than expected growth in sales translates to a 0.53 percentage point increase in gross investment.

Table 7: Does the Price Setting Type Matter?

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
Time-dep. Pricing \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.16 (0.16)	0.03*** (0.01)
State-dep. Pricing \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.52*** (0.15)	0.05*** (0.01)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.08*** (0.14)	0.07*** (0.01)
$\Delta\text{CAPEX}_{t-4 t}$	-0.26*** (0.02)	0.00 (0.00)
$\Delta\text{Price}_{t-4 t}$	-0.05 (0.41)	0.39*** (0.03)
Time-dep. = State-dep. (p-value)	0.061	0.141
N	2,580	2,303

Note: This table reports the results from estimating model (5), controlling for dummies for the respective interaction terms (omitted for brevity). Time-dependent pricing refers to firms that reset their price in regular intervals, whereas state-dependent Pricing refers to firms who reset their prices depending on the state of their business. Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Column (2) reports the estimated price response to sales growth surprises for both types of firms. As for CAPEX, state-dependent pricing firms exhibit a stronger response to sales growth surprises than time-dependent ones. Specifically, for more attentive, state-dependent firms a 1 percentage point higher than expected sales growth generates a 0.05 percentage

point increase in prices.¹⁷ Importantly, the price response is calculated four-quarters after the realisation of the sales growth surprise, an horizon over which the vast majority of time-dependent pricing firms also resets prices. This implies that the significant response of prices for time-dependent firms is likely the cumulative effect of all additional shocks that these firms experience in between their price setting schedules.

Figure C.1 shows that the distribution of realised sales growth surprises is very similar for both types of price-setters: Both distributions are centered around zero, but the distribution of sales growth surprises for state-dependent price-setters has less mass around the center and is slightly more skewed towards the left. We have shown in Table C.5 that firms react to large, positive sales growth surprises in particular. Given that these occur relatively more frequently to time-dependent price setters than to state dependent-price setters, the stronger response of state-dependent price setters is even more striking.

5.2 Alternative Explanations

In this section, we explore four alternative drivers behind our results. First, we examine financial frictions as a potential driver of the observed MPI. Second, we investigate the dynamics of uncertainty as a fundamental determinant of firms' sales growth surprises. Third, we analyze the link between sales growth surprises and idiosyncratic productivity shocks. Finally, we assess whether firms respond differently to micro and macro shocks.

5.2.1 Financial Constraints

An alternative explanation for the presence of a positive MPI on average is that financially constrained firms use unexpected income to invest because they cannot obtain these funds externally, resulting in a positive correlation between sales surprises and gross investments. If financial constraints were the main driver of firms investment response we would expect MPIs to be different for firms' experiencing different levels of financial constraints. In particular, we test the role of financial constraints by estimating MPIs (using Equation 5) across

¹⁷These results remain unchanged when also interacting sales growth surprises with the price-setting type dummy to account for a potentially heterogeneous updating of expectations between the two groups.

different firm characteristics - specifically age, size and balance sheet measures of firms' financial positions - that have been frequently used as proxies for financial constraints (see [Ottonello and Winberry, 2020](#); [Cloyne et al., 2023](#); [Jeenas, 2023](#)).

Age and size. The first characteristics that we use as proxies for financial constraints are firms' age and size.¹⁸ Intuitively, as firms grow, on average they become less financially constrained. If financial constraints were to be a significant driver of MPIs we would expect the coefficients to be decreasing in firm age and size. We classify firms as young if they are less than 10 years old, medium if they are between 10 and 20 years old, and old if they are more than 20 years old. We classify firms as small if they have fewer than 50 employees, medium if they have between 50 and 249 employees, and large if they have 250 or more employees (see Table B.7 for summary statistics of the sales growth surprises for each group).

Table 8 reports the interaction terms between the sales growth surprises and the three categories of firm age, Column (1), and size in Column (2). In both cases, the coefficients indicate that firms in the middle of the distribution respond strongest to sales growth surprises. However, for none of the three groups we can reject the null hypothesis that they react significantly different from the other groups. As an extension, we also interact firm age and size, grouping firms into four categories: young and small, young and medium & large, old and small, old and medium/large. Table C.12 shows that, if anything, larger and older firms react more strongly to sales growth surprises.

Balance sheet measures. To complement the previous analysis, we also test whether MPIs exhibit any significant heterogeneity across a set of balance sheet measures that capture more directly firms' financial positions. We focus on three main indicators: leverage, net liquidity, and the interest burden, measured as the ratio of interest expenditure to sales.¹⁹ Table B.6 reports the median firm characteristic for each tercile.

¹⁸[Cloyne et al. \(2023\)](#) and [Jeenas \(2023\)](#) show how age and size are good proxies for firms' financial constraints.

¹⁹For each of these variables we assign firms to terciles of the respective distribution in each quarter based on their latest available balance sheet information to. This implies that the respective balance sheet variable can be up to 3 quarters old.

Table 8: The Role of Financial Constraints I - Age and Size

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se
Young \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.42 (0.35)	
Medium \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.31** (0.15)	
Old \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.29*** (0.11)	
Small \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.02 (0.20)
Medium \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.46*** (0.12)
Large \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.18 (0.15)
1st = 2nd (p-value)	0.77	0.06
2nd = 3rd (p-value)	0.89	0.15
N	4,271	4,001

Note: This table reports the results from estimating model (5), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). We classify firms as young if they are less than 10 years old, medium if they are between 10 and 20 years old, and old if they are more than 20 years old. We classify firms as small if they have fewer than 50 employees, medium if they have between 50 and 249 employees, and large if they have 250 or more employees. Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. See Tables B.6 & B.7 for the respective median value of the firm characteristic and median sales growth surprises for each group.

Table C.13 reports the main coefficients of interest. Column (1) shows MPIs for different terciles of the leverage distribution. The results show how MPIs are significant only for firms with low and medium leverage, although it is not possible to statistically distinguish the responses across the leverage terciles.

Column (2), instead, reports MPIs for different terciles of liquidity. Interestingly, unexpected sales realisations do not transmit differently to gross investment for firms with different liquidity position. The measured MPIs are remarkably stable for all terciles of the liquidity distribution, indicating that 1 percentage point higher than expected sales growth induces

approximately 0.3 percentage points more investment for both firms in the top and bottom terciles of the liquidity distribution.

Finally, Column (3) shows MPI estimates across different terciles of the interest burden distribution. Firms with a high interest burden should benefit more from the windfall generated by higher than expected sales growth. However, we cannot reject the hypothesis that the estimated MPIs differ significantly across different terciles of the interest burden distribution.

Profitability. Finally, we investigate profitability as an indicator of firms’ financial constraints. This hypothesis implies that highly profitable firms face less severe constraints in investment decisions compared to their less profitable counterparts. For this, more profitable firms, characterised by higher margins, and those with elevated markups compared to the industry average, should exhibit less sensitivity to sales growth surprises. However, we find that firms with higher margins and higher markups (relative to the industry average) tend to adjust their investment more strongly in response to sales growth surprises, as shown in Table C.14 and C.15.²⁰ This is consistent with the idea that market power provides a strong incentive to pay attention to the evolution of firms’ own demand.

Taken together, these results indicate that the role of financial constraints in shaping firms’ marginal propensity to invest is, if anything, small. Furthermore, if the estimated marginal propensity to invest would predominantly reflect the role of financial constraints, we would not expect a significant response of prices.

5.2.2 Uncertainty

The structure of the DMP questions allows us to construct a unique measure of firm-level uncertainty: Firms participating in the survey are asked to report a distribution for their sales expectations with relative probability weights.

²⁰To compute relative mark-ups, we use the expression for the markup $\mu_{it} = \theta_{it}^v \frac{P_{it}Q_{it}}{P_{it}^v V_{it}}$ of De Loecker, Eeckhout and Unger (2020) where θ_{it}^v is the industry-specific output elasticity, and $\frac{P_{it}Q_{it}}{P_{it}^v V_{it}}$ the revenue share of the variable input. Taking logs and demeaning this expression on the industry-quarter level eliminates the industry-specific constant and thus returns the markup of the firm relative to its respective industry. We compute $\frac{P_{it}Q_{it}}{P_{it}^v V_{it}}$ as the sales-to-cost-of-goods-sold ratio.

This allows us to test two hypothesis: First, it might be the case that sales growth surprises lead to higher sales growth uncertainty going forward, which would affect firm investment via the real options channel. In other words, sales growth surprises might affect investment because they proxy changes in uncertainty. To test this, we include sales uncertainty and its four-periods lag in Equation (2). Second, if firms extract information about the state of their demand functions from the realisation of sales growth surprises, we should expect a positive relationship between MPIs and the level of uncertainty about firms' sales forecasts. In other words, firms that are more uncertain about their future sales should extract a large signal from different sales realisations. To test this, we estimate Equation (5), interacting sales growth surprises with dummies indicating the terciles of the firm-level uncertainty distribution in the DMP sample.

Table 9 displays the respective results. Column (1) shows that controlling for sales uncertainty hardly changes our baseline estimate suggesting that sales growth surprises affect investment not just because they change sales growth uncertainty. In Column (2) we test whether firms that are more uncertain about their future sales do react more to unexpected sales growth realisations. The coefficients are significant only for firms in the middle and top terciles of the uncertainty distribution. Here, we cannot fully reject the hypothesis that firms with high uncertainty react differently than firms with low uncertainty although the p-value is close to usual significance thresholds.

As a robustness check, we replace the standard deviation of sales growth as measure of subjective uncertainty with the span between the highest and lowest sales realisations expected by firms (as in [Bachmann et al. \(2021\)](#)). Column (1) of Table C.16 shows that the estimated average MPI remains unchanged. However, we cannot reject the null hypothesis of equal responsiveness between different terciles of the uncertainty distribution.

5.2.3 Productivity

Another potential channel that would generate a positive relationship between sales growth surprises and investments are firm-level productivity shocks. As firm productivity improves

Table 9: The Role of Sales Uncertainty

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.31*** (0.08)	
Low Uncertainty $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.00 (0.17)
Medium Uncertainty $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.34** (0.13)
High Uncertainty $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.41*** (0.13)
Sales Uncertainty	0.40 (0.46)	
L4.Sales Uncertainty	-0.28 (0.41)	
1st = 2nd (p-value)		0.12
2nd = 3rd (p-value)		0.68
N	4,312	4,312

Note: This table reports the results from estimating models (2) (Column (1)) as well (5) (Column (2)), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). Uncertainty refers to the standard deviation of sales growth expectations. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. See Tables B.6 & B.7 for the respective median value of the firm characteristic and median sales growth surprises for each group.

unexpectedly, firms can sell more goods at the same price. As a consequence, firms optimally respond by increasing their factor demands, thus increasing their gross investment.

To test this channel we augment our baseline specification (Equation (2)) with controls for changes in firm-level total factor productivity (TFP) and labour productivity.²¹ If the effect

²¹We estimate TFP as

$$\ln(\text{TFP}_{i,t}) = \ln(\text{GVA}_{i,t}) - 0.63 * \ln(\text{Labour Costs}_{i,t}) - 0.37\ln(\text{Fixed Assets}_{i,t})$$

and labour productivity as

$$\ln(\text{Labour Productivity}_{i,t}) = \ln \frac{\text{GVA}_{i,t}}{\# \text{ of Employees}_{i,t}}$$

where gross value added (GVA) is the sum of operating profits and labour costs. All values are deflated using the GVA deflator.

of sales surprise on CAPEX were only due to these firm specific productivity shocks, adding them as controls should render the coefficient on the sales growth surprises insignificant.

Table 10: The Role of Productivity

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.43*** (0.11)	0.41*** (0.11)
$\Delta\text{TFP}_{t-4 t}$	6.62 (7.27)	
$\Delta\text{Labour Prod.}_{t-4 t}$		7.32 (7.00)
Covid	Yes	Yes
FE	No	No
N	2,801	2,823

Note: This table reports the results from estimating model (2), controlling for realised CAPEX growth, realised price growth, expected sales growth (omitted for brevity). For details on the computation of productivity, see Footnote 21. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

However, as shown in Table 10, controlling for the change in either firm specific TFP or labor productivity does not significantly change the estimate of the MPI, and in both cases the productivity measures remain insignificant. If firms' sales growth surprises were predominantly a proxy for productivity shocks, we would instead expect the coefficient on the sales growth surprises to turn insignificant.

On top of this, it is unlikely that sales growth surprises simply proxy for productivity shocks because that would be at odds with the positive response of prices to sales growth surprises: if firms' sales growth surprises would predominantly capture productivity shocks, we would expect a price response with a different sign. In a standard setting in which firms' enjoy some constant degree of market power and prices are relatively flexible, a positive productivity shock would in fact decrease firms' marginal costs, leading to a decrease in prices rather than an increase. We take this as evidence that our measure of sales surprises is not capturing firm specific productivity shocks.

5.2.4 Micro vs Macro Shocks

In our baseline specification we are agnostic about the source of the sales growth surprises. In light of recent evidence by [Born et al. \(2022\)](#), we also test whether firms react differently to micro and macro shocks. To do so, we first decompose sales growth surprises into a micro and a macro component by regressing them on industry \times quarter dummies. We interpret the fitted values (residuals) of this regression as macro (micro) shocks. To investigate the potentially heterogeneous response, we standardise the two respective shocks and include them separately in our baseline model (2). The results in Table C.17 indicate that firms do not adjust their gross investment in a way that is statically significantly different in response to micro or macro news.²² However, firms increase their prices significantly more in response to sales growth surprises that are driven by macro shocks compared to sales growth surprises driven micro shocks.

6 Conclusion

This study introduces a novel approach to calculate the Marginal Propensity to Invest of UK firms, leveraging unique firm-level data from the Decision Maker Panel survey and firm balance sheet data from Bureau van Dijk. Our methodology constructs income shocks by utilizing survey-based sales forecast errors in a panel of UK firms, allowing for a detailed analysis of MPI dynamics. Our analysis reveals a significantly positive relationship between unexpected sales growth and gross investment, indicating that firms adjust their investment in response to changes in demand conditions. We find that approximately one-third of an unexpected increase in sales is translated into investments, a magnitude greater than previously observed in literature. Our results suggest a behavioral response, where firms adjust investments based on learning from income shocks, particularly more attentive firms. We find limited support in the data for alternative explanations, such as financial constraints,

²²Decomposing the sales growth surprises into idiosyncratic and narrower two-digits sectors (instead of one-digit) indicates that firms appear to react more to the narrower industry component than the idiosyncratic one. However, it is unclear how the narrower industry definition is able to distinguish between macro and micro news versus other potential channels, like strategic complementarities in response to aggregate shocks.

productivity, and uncertainty. Additionally, we estimate the price response to income shocks, finding a significant increase in prices following unexpected sales growth. This study contributes to understanding how fluctuations in firms' income affect their investment decisions. They also provide valuable insights into the UK economic landscape by shedding light on the behavioral aspects of UK firms' investment strategies.

References

- Altig, Dave, Scott Baker, Jose Maria Barrero, Nicholas Bloom, Philip Bunn, Scarlet Chen, Steven J. Davis, Julia Leather, Brent Meyer, Emil Mihaylov, Paul Mizen, Nicholas Parker, Thomas Renault, Pawel Smietanka, and Gregory Thwaites. 2020. “Economic uncertainty before and during the COVID-19 pandemic.” *Journal of Public Economics*, 191: 104274.
- Bachmann, Ruediger, Kai Carstensen, Stefan Lautenbacher, and Martin Schneider. 2021. “Uncertainty and Change: Survey Evidence of Firms’ Subjective Beliefs.” National Bureau of Economic Research Working Paper, National Bureau of Economic Research.
- Bahaj, Saleem, Angus Foulis, and Gabor Pinter. 2020. “Home values and firm behavior.” *American Economic Review*, 110(7): 2225–2270.
- Barrero, Jose Maria. 2022. “The Micro and Macro of Managerial Beliefs.” *The Journal of Financial Economics*, 143(2): 640–667.
- Baumeister, Christiane, and James D Hamilton. 2023. “A Full-Information Approach to Granular Instrumental Variables.” Working Paper, UCSD.
- Berman, Nicolas, Vincent Rebeyrol, and Vincent Vicard. 2019. “Demand Learning and Firm Dynamics: Evidence From Exporters.” *Review of Economics & Statistics*, 101(1): 91–106.
- Bloom, Nicholas. 2009. “The Impact of Uncertainty Shocks.” *Econometrica*, 77(3): 623–685.
- Bloom, Nicholas, Philip Bunn, Scarlet Chen, Paul Mizen, Pawel Smietanka, and Gregory Thwaites. 2019. “The impact of Brexit on UK firms.” National Bureau of Economic Research.
- Bloom, Nicholas, Steven J. Davis, Lucia S. Foster, Scott W. Ohlmacher, and Itay Saporta-Eksten. 2022. “Investment and subjective uncertainty.” NBER working paper no. 30654.
- Bloom, Nick, Stephen Bond, and John Van Reenen. 2007. “Uncertainty and investment dynamics.” *The Review of Economic Studies*, 74(2): 391–415.
- Blundell, Richard, Luigi Pistaferri, and Ian Preston. 2008. “Consumption Inequality and Partial Insurance.” *American Economic Review*, 98(5): 1887–1921.
- Born, Benjamin, Zeno Enders, Manuel Menkhoff, Gernot J Müller, and Knut Niemann. 2022. “Firm expectations and news: micro v macro.” CESifo, CESifo Working Paper.
- Boutros, Michael, Itzhak Ben-David, John R Graham, Campbell R Harvey, and John W Payne. 2020. “The persistence of miscalibration.” National Bureau of Economic Research.
- Cloyne, James, Clodomiro Ferreira, Maren Froemel, and Paolo Surico. 2023. “Monetary Policy, Corporate Finance, and Investment.” *Journal of the European Economic Association*, 21(6): 2586–2634.
- Cummins, Jason G, Kevin A Hassett, and R Glenn Hubbard. 1994. “A reconsideration of investment behavior using tax reforms as natural experiments.” *Brookings papers on economic activity*, 1994(2): 1–74.
- Davis, Steven J, John C Haltiwanger, and Scott Schuh. 1998. “Job creation and destruction.” *MIT Press Books*, 1.
- De Loecker, Jan, Jan Eeckhout, and Gabriel Unger. 2020. “The rise of market power and the macroeconomic implications.” *The Quarterly Journal of Economics*, 135(2): 561–644.
- Evans, George W, and Seppo Honkapohja. 2001. *Learning and expectations in macroeconomics*. Princeton University Press.
- Hebous, Shafik, and Tom Zimmermann. 2021. “Can government demand stimulate private

- investment? Evidence from US federal procurement.” *Journal of Monetary Economics*, 118: 178–194.
- Jeenas, Priit.** 2019. “Monetary Policy Shocks, Financial Structure, and Firm Activity: A Panel Approach.” mimeo.
- Jeenas, Priit.** 2023. “Firm Balance Sheet Liquidity, Monetary Policy Shocks, and Investment Dynamics.”
- Kaplan, Steven N, and Luigi Zingales.** 1997. “Do investment-cash flow sensitivities provide useful measures of financing constraints?” *The Quarterly Journal of Economics*, 112(1): 169–215.
- Khan, Aubhik, and Julia K. Thomas.** 2013. “Credit Shocks and Aggregate Fluctuations in an Economy with Production Heterogeneity.” *Journal of Political Economy*, 121(6): 1055–1107.
- Lakdawala, Aeimit, and Timothy Moreland.** 2022. “Firm-Level Uncertainty and the Transmission of Monetary Policy.” Available at SSRN: <https://ssrn.com/abstract=4255980> or <http://dx.doi.org/10.2139/ssrn.4255980>.
- Martin-Baillon, Alaïs.** 2021. “Firms’ Marginal Propensities to Invest.”
- Nickell, Stephen.** 1981. “Biases in Dynamic Models with Fixed Effects.” *Econometrica*, 49(6): 1417–1426.
- Ottanello, Pablo, and Thomas Winberry.** 2020. “Financial heterogeneity and the investment channel of monetary policy.” *Econometrica*, 88(6): 2473–2502.

APPENDIX

A Questions

To illustrate the survey questions asked in the DMP, we list the respective questions asked for investment, sales, and prices in 2016q4.

A.1 Investment

Looking backwards, firms were asked about their investment (in thousand £) in the previous quarter and the corresponding quarter one year before. Realised investment growth is then computed as the growth rate between these two values.

"In the last quarter (July - September 2016), what was the approximate sterling value of your CAPITAL EXPENDITURE (in £, THOUSANDS)?"

"Looking back over the past year, what was the approximate sterling value of your CAPITAL EXPENDITURE in the same quarter a year ago (July - September 2015) (in £, THOUSANDS)?"

Looking forward, firms were asked about their investment (in thousand £) 3 quarters ahead in 5 different scenarios and the corresponding probabilities. Expected investment growth is then computed as the growth rate between the probability weighted level of investment three quarters ahead and last quarter's investment.

"Looking a year ahead from the last quarter (July - September 2016), what would be the approximate sterling value of CAPITAL EXPENDITURE you expect for the same quarter (July - September 2017) in each of the following scenarios?" (with five scenarios provided; i) lowest, ii) low, iii) middle, iv) high, v) highest)

"Please assign a percentage likelihood (probability) to the amounts of CAPITAL EXPENDITURE you entered."

A.2 Sales

Looking backward, firms were asked about the growth of their sales revenue between the previous quarter and the corresponding quarter one year before.

"Looking back over the past year from the third quarter of 2016 (July - September), by what % amount has your SALES REVENUE changed since the same quarter a year ago (July - September 2015)?"

Looking forward, firms were asked about the growth rate of their sales between the previous quarter and the quarter 3 periods ahead in 5 different scenarios and the corresponding probabilities. Expected sales growth is then computed as the probability weighted sales growth.

”Looking a year ahead from the last quarter (July - September 2016), by what % amount do you expect your SALES REVENUE to have changed in each of the following scenarios?” (with five scenarios provided; i) lowest, ii) low, iii) middle, iv) high, v) highest)”

”Please assign a percentage likelihood (probability) to the % changes in SALES REVENUE you entered (values should sum to 100%).”

A.3 Prices

Looking backward, firms were asked about the growth of their average prices between the current month and the corresponding month one year before.

”Looking back, from 12 months ago to now, what was the approximate % change in the AVERAGE PRICE you charge, considering all products and services?”

Looking forward, firms were asked about the growth of their average prices between the current month and the month one year ahead in 5 different scenarios and the corresponding probabilities. Expected price growth is then computed as the probability weighted price growth.

”Looking ahead, from now to 12 months from now, what approximate % change in your AVERAGE PRICE would you assign to each of the following scenarios?” (with five scenarios: lowest, low, middle, high, highest provided)

” Please assign a percentage likelihood (probability) to the % changes in your AVERAGE PRICES you entered.”

B Data

Table B.1: Summary Statistics - DMP

(1)						
	N	Mean	s.d.	p10	p50	p90
$\mathbb{E}_{t+1}\Delta\text{Sales}_{t t+4}$	53,983	7.32	10.76	-3.00	5.00	21.45
$\Delta\text{Sales}_{t-4 t}$	55,993	6.62	20.89	-19.00	5.00	34.00
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	21,768	-0.72	20.39	-24.79	-0.20	23.00
$\sigma(\Delta\text{Sales}_{t t+4})$	53,983	5.90	4.49	1.49	4.66	12.77
$\mathbb{E}_t\Delta\text{Price}_{t t+4}$	44,649	3.49	3.27	0.00	2.90	7.75
$\Delta\text{Price}_{t-4 t}$	46,127	3.93	4.63	0.00	3.00	10.00
$\mathbb{E}_{t+1}\Delta\text{CAPEX}_{t t+4}$	48,341	10.58	81.09	-101.53	4.08	132.20
$\Delta\text{CAPEX}_{t-4 t}$	46,498	3.94	88.40	-127.27	0.00	133.33
Learning Gain γ_i	12,004	-0.09	0.29	-0.48	-0.06	0.25

Table B.2: Summary Statistics - BvD

(1)						
	N	Mean	s.d.	p10	p50	p90
Age	18,315	29.00	23.09	8.00	22.00	59.00
Employees	21,103	748.48	7930.81	17.00	94.00	664.00
Total Assets	21,823	8.2e+05	2.7e+07	1798.10	11998.00	1.6e+05
Leverage	20,978	50.46	31.22	11.80	47.48	89.79
Liquidity	20,899	-2061.96	2.3e+05	-12.10	22.64	64.29
Interest Burden	13,360	2.23	6.95	0.02	0.39	4.29
$\Delta_4 \ln(\text{Labour Productivity})$	10,098	0.00	0.45	-0.35	0.00	0.36
$\Delta_4 \ln(\text{TFP})$	9,999	-0.01	0.42	-0.39	0.00	0.34

Note: This table reports the summary statistics of variables obtained from BvD for the years 2016-2022. All variables (except age) are winsorised by reporting quarter at the 5th and 95th percentile. Net liquidity is computed as the ratio of liquid assets minus short-term obligations to total assets. Interest burden is computed as the ratio of interest payments to total sales. For details on the computation of productivity, see Footnote 21.

Table B.3: Predictability of Sales Growth Surprises

	(1)	(2)	(3)	(4)
	$FE_{t+1}\Delta Sales_{t t+4}$	$FE_{t+1}\Delta Sales_{t t+4}$	$FE_{t+1}\Delta Sales_{t t+4}$	$(FE_{t+1}\Delta Sales_{t t+4})$
	b/se	b/se	b/se	b/se
$\Delta Sales_{t-4 t}$	-0.02 (0.02)			
$L4.E_{t+1}\Delta Sales_{t t+4}$		-0.09* (0.05)		
$FE_{t-3}(\Delta Sales_{t-4 t})$			-0.00 (0.02)	
$(FE_{t-3}\Delta Sales_{t-4 t})$				0.18*** (0.03)
N	3,746	3,746	3,746	3,746

Note: This table reports the relationship between sales forecast errors and the preceding realised sales growth (Column (1)), sales forecast error (Column (2)), and absolute sales forecast error (Column (3)) in the sample used to estimate model (2). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table B.4: Forecast Revisions and Sales Growth Surprises

	(1)	(2)
	$FE_{t+1}\Delta Sales_{t t+4}$	$FE_{t+1}\Delta Sales_{t t+4}$
	b/se	b/se
$E_{t+1}\Delta Sales_{t+4 t} - E_t\Delta Sales_{t+3 t-1}$	0.28*** (0.04)	
Micro News		3.41*** (0.52)
Macro & Industry News		-0.57 (0.74)
Micro = Macro (p-value)		0.00
N	3,320	3,320

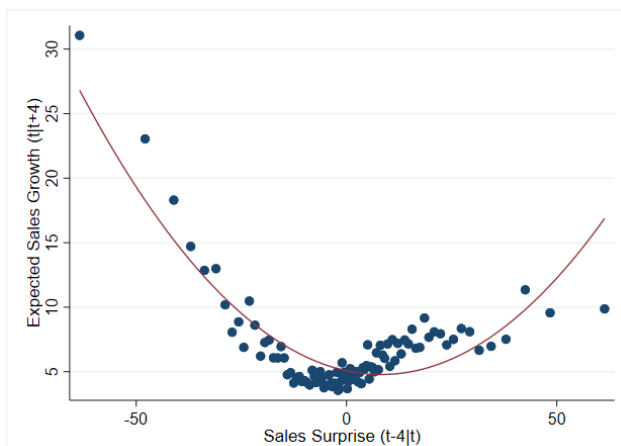
Note: This table reports the relationship between sales forecast errors and the preceding sales growth (Column (1)), sales forecast revision (Column (2)), and sales forecast revision decomposed into a common component within sector and an idiosyncratic one (Column (3)) (2). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table B.5: Long-run Effects of Sales Surprises

	(1)	(2)
	$\mathbb{E}_{t+1}\Delta\text{Sales}_{t t+4}$	$\mathbb{E}_{t+5}\Delta\text{Sales}_{t+4 t+8}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})^-$	-0.44***	-0.14***
	(0.06)	(0.02)
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})^+$	0.07***	-0.01
	(0.03)	(0.02)
Negative = Positive (p-value)	0.00	0.00
N	4,312	3,770

Note: This table reports the relationship between positive and negative sales forecast errors and the subsequent expectations for sales growth over the period $(t, t + 4)$ as well as $(t + 4, t + 8)$ (Column (2)) in the sample used to estimate model (2). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Figure B.1: Sales Growth Surprises and Expectation Revisions



Note: This figure displays the binned distribution of sales growth surprises and subsequent sales growth expectations.

Table B.6: Group Median of Sample Splits

	Median 1 st Group	Median 2 nd Group	Median 3 rd Group
Age	7	15	33
Employees	26	99	439
Leverage	18.67	47.06	79.73
Net Liquidity	-1.39	23.14	55.30
Interest Burden	.08	.64	4.36
Uncertainty	2.05	4.71	10.07
Gain	-.38	-.02	.32

Note: This table reports the median value for each respective variable used in the dummy interactions terms of model (5).

Table B.7: Summary of Sales Growth Surprise by Sample Split

	1 st Group	2 nd Group	3 rd Group
Age	-2.13 (21.27)	-0.96 (20.03)	-0.38 (20.32)
Employees	-1.55 (22.67)	-0.74 (20.36)	0.06 (17.58)
Leverage	-1.39 (21.06)	-0.40 (20.85)	-0.59 (21.39)
Liquidity	-1.31 (21.14)	-0.46 (20.92)	-0.89 (21.41)
Interest Burden	1.09 (19.96)	0.41 (19.54)	-2.61 (21.48)
Gain	-1.55 (18.86)	0.74 (19.05)	1.19 (17.53)
Sales Uncertainty	0.72 (13.64)	-0.08 (18.54)	-1.76 (25.83)

Note: This table reports the mean sales growth surprise for each respective variable category used in the dummy interactions terms of model (5).

C Further Results

C.1 Main Results: Sensitivity and Extensions

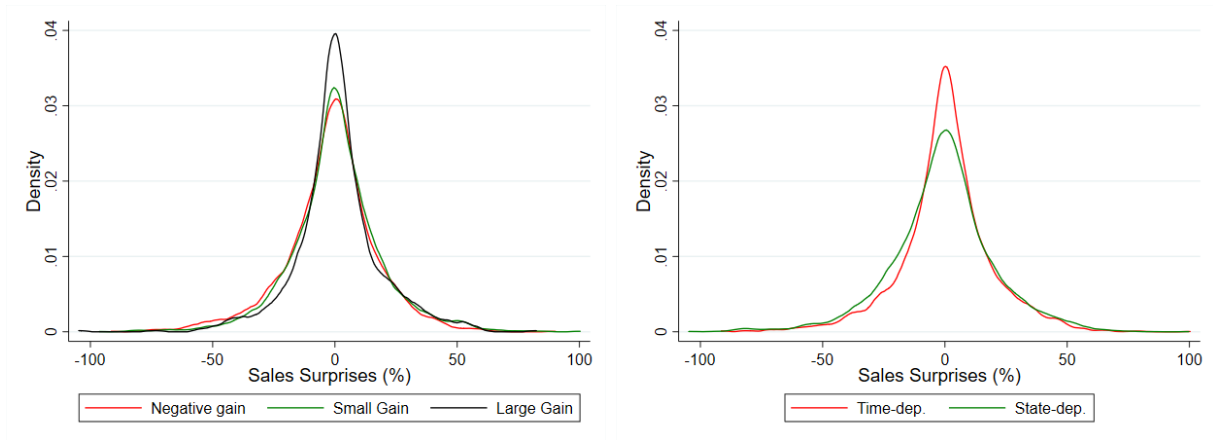
Table C.1: Average MPI - Not Controlling for Expectations

	(1)	(2)	(3)	(4)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.21** (0.10)	0.38*** (0.09)	0.22** (0.09)	
$\text{FE}_{t-3}(\text{med}\Delta\text{Sales}_{t-4 t})$				0.21** (0.10)
$\Delta\text{CAPEX}_{t-4 t}$	-0.25*** (0.02)	-0.24*** (0.02)	-0.25*** (0.02)	-0.25*** (0.02)
$\Delta\text{Price}_{t-4 t}$	0.01 (0.38)	0.19 (0.39)	-0.52 (0.38)	0.00 (0.38)
Covid	Yes	No	Yes	Yes
FE	No	No	Quarter x Ind.	No
N	4,312	3,918	4,290	4,312

Note: This table reports the results from estimating model (2) omitting the sales growth expectations control. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

C.2 Drivers: Results and Extensions

Figure C.1: Density of Sales Growth Surprises by learning gain and price-setting type



(a) Magnitude of Learning Gain

(b) Price-setting Type

Note: Panel a of this figure plots the density of realised sales growth surprises for firms with large negative gains (red), small gains (green), and large positive gains (black). Panel b of this figure plots the density of realised sales growth surprises for time-dependent price setters (red) and state-dependent price-setters (green)

Table C.2: Average MPI - Balance Sheet Data

	(1)	(2)	(3)
	$\Delta\text{Fix. Assets}_{t t+4}$	$\Delta\text{Tang. Assets}_{t t+4}$	$\Delta\text{Intan. Assets}_{t t+4}$
	b/se	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.17*** (0.05)	0.16** (0.07)	0.06 (0.17)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	0.27*** (0.08)	0.28*** (0.10)	0.60*** (0.19)
$\Delta\text{Price}_{t-4 t}$	0.25 (0.32)	0.16 (0.49)	-0.73 (0.80)
$\Delta\text{CAPEX}_{t-4 t}$			
$\Delta\text{Fix. Assets}_{t-4 t}$	0.03 (0.04)		
$\Delta\text{Tang. Assets}_{t-4 t}$		0.00 (0.02)	
$\Delta\text{Intan. Assets}_{t-4 t}$			0.07* (0.04)
Covid	Yes	Yes	Yes
FE	No	No	No
N	1,291	1,246	536

Note: This table reports the results from estimating model (2), replacing the growth rate of CAPEX (as reported in the DMP) as dependent variables with the growth rate of the fixed asset stock (Column (1)), the tangible asset stock (Column (2)), and the intangible asset stock (Column (3)) as reported in BvD. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.3: Average MPI - Medium-term Response

	(1)	(2)
	$\Delta\text{CAPEX}_{t+4 t+8}$	$\Delta\text{Prices}_{t+4 t+8}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	-0.20*	0.00
	(0.10)	(0.01)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	-0.16	0.06***
	(0.20)	(0.01)
$\Delta\text{CAPEX}_{t-4 t}$	0.02	-0.00**
	(0.03)	(0.00)
$\Delta\text{Price}_{t-4 t}$	-0.44	-0.01
	(0.70)	(0.05)
Covid	Yes	Yes
FE	No	No
N	2,036	2,457

Note: This table reports the results from estimating model (2), using two-year ahead investment & price growth as dependent variable. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.4: Average MPI - Quantity Surprises

	(1)	(2)
	$\Delta\text{Real CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Quantity}_{t-4 t})$	0.35***	0.03***
	(0.12)	(0.01)
$\mathbb{E}_{t+1}(\Delta\text{Quantity}_{t t+4})$	1.13***	0.03***
	(0.14)	(0.01)
$\Delta\text{Real CAPEX}_{t-4 t}$	-0.28***	0.00
	(0.02)	(0.00)
$\Delta\text{Price}_{t-4 t}$	-0.28	0.48***
	(0.62)	(0.03)
Covid	Yes	Yes
FE	No	No
N	2,314	2,729

Note: This table reports the results from estimating model (2), deflating CAPEX growth using the CPI and deflating sales growth (expectations) with the firm's own price growth (expectations). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.5: Magnitude of Sales Surprises

	(1)
	$\Delta\text{CAPEX}_{t t+4}$
	b/se
Lower Tercile of $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	-0.26 (0.20)
Medium Tercile of $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.57 (0.62)
Upper Tercile of $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.69*** (0.19)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	0.88*** (0.12)
$\Delta\text{CAPEX}_{t-4 t}$	-0.25*** (0.02)
$\Delta\text{Price}_{t-4 t}$	-0.11 (0.37)
Small = Medium (p-value)	0.22
Medium = Large (p-value)	0.85
N	4,312

Note: This table reports the results from estimating model (2), interacting sales growth surprises with dummies indicating in which tercile of the distribution the sales growth surprises falls. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.6: Response by Sector

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
Manufacturing & Construction $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.23** (0.11)	0.04*** (0.01)
Services $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.34*** (0.12)	0.03*** (0.01)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.00*** (0.12)	0.07*** (0.01)
$\Delta\text{CAPEX}_{t-4 t}$	-0.25*** (0.02)	0.00 (0.00)
$\Delta\text{Price}_{t-4 t}$	-0.12 (0.38)	0.45*** (0.03)
Manufacturing & Construction = Services (p-value)	0.49	0.69
N	4,249	3,651

Note: This table reports the results from estimating model (5), controlling for dummies for the respective interaction term (omitted for brevity). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.7: Average MPI - Controlling for Survey Features

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.22*	0.03***
	(0.12)	(0.01)
$\text{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.12***	0.05***
	(0.14)	(0.01)
$\Delta\text{CAPEX}_{t-4 t}$	-0.26***	-0.00
	(0.02)	(0.00)
$\Delta\text{Price}_{t-4 t}$	0.45	0.36***
	(0.68)	(0.05)
Duration of Survey Participation	6.55***	0.72***
	(1.60)	(0.09)
Response by CFO	2.65	-0.36
	(5.85)	(0.31)
Response by Other	0.15	-0.48
	(10.14)	(0.53)
Covid	Yes	Yes
FE	No	No
N	2,299	1,991

Note: This table reports the results from estimating model (2), controlling for additional survey features as well as dummies for the legal form for the firm (omitted for brevity). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.8: Employment Weights and IV

	(1)	(2)	(3)	(4)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.32** (0.14)		0.03*** (0.01)	
$\Delta\text{Sales}_{t-4 t}$		0.33*** (0.09)		0.04*** (0.01)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.05*** (0.27)	0.95*** (0.12)	0.04** (0.02)	0.06*** (0.01)
$\Delta\text{Price}_{t-4 t}$	0.05 (1.06)	-0.23 (0.38)	0.50*** (0.06)	0.44*** (0.03)
$\Delta\text{CAPEX}_{t-4 t}$	-0.23*** (0.03)	-0.26*** (0.02)	-0.00* (0.00)	-0.00 (0.00)
Covid	Yes	Yes	Yes	Yes
FE	No	No	No	No
N	4001.00	4312.00	3473.00	3704.00
Weights	Yes	No	Yes	No
IV	No	Yes	No	Yes
F-statistic		70.82		107.17

Note: This table reports the results from estimating model (2) using employment weights (columns (1) and (3)) and using the sales growth surprises as instrument for sales growth expectations (columns (2) and (4)). Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.9: Descriptive Statistics by Size of Learning Gain

	$\gamma \in (-1, -1/5)$	$\gamma \in (-1/5, 1/5)$	$\gamma \in (1/5, 1)$
Age	35.63 (27.66)	34.79 (27.51)	31.02 (22.17)
Sales	94,576.26 (639393.79)	240249.96 (9.65e+06)	218544.41 (1.37e+06)
Employees	533.00 (1,221.19)	512.82 (1,321.66)	368.66 (817.77)
Sales Volatility	5.23 (7.28)	6.27 (7.59)	8.10 (9.61)
Sales Uncertainty	5.40 (2.65)	4.97 (3.07)	5.07 (2.89)
Manufacturing & Construction (%)	0.30 (0.46)	0.33 (0.47)	0.24 (0.43)

Note: This table reports the mean summary statistic of firms by the magnitude of their respective learning gain. Standard errors of the respective value are reported in parenthesis.

Table C.10: Persistence by Attentiveness

	(1)	(2)
	$\Delta\text{Sales}_{t t+4}$	$\Delta\text{Sales}_{t t+4}$
	b/se	b/se
$\gamma_i \in (-1, -1/5) \times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	-0.03 (0.04)	
$\gamma_i \in (-1/5, 1/5) \times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	-0.04 (0.03)	
$\gamma_i \in (1/5, 1) \times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.06 (0.09)	
Time-dep. Pricing $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		-0.01 (0.03)
State-dep. Pricing $\times \text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.00 (0.04)
$\Delta\text{Price}_{t-4 t}$	-0.13 (0.10)	-0.20* (0.10)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	0.73*** (0.06)	0.63*** (0.06)
Low gain = Medium gain (p-value)	0.81	
Medium gain = High gain (p-value)	0.27	
Low gain = High gain (p-value)	0.36	
Time-dep. = State-dep. (p-value)		0.85
N	2,567	2,303

Note:

Table C.11: Descriptive Statistics by Price-setting Strategy

	Time-dep.	State-dep.
Age	30.84 (24.57)	32.50 (25.11)
Sales	221800.70 (9.39e+06)	190332.71 (7.39e+06)
Employees	478.36 (3,005.85)	376.44 (1,505.74)
Sales Volatility	8.10 (11.47)	6.23 (11.91)
Sales Uncertainty	5.45 (3.67)	6.33 (3.88)
Manufacturing & Construction (%)	0.21 (0.40)	0.32 (0.46)

Note: This table reports the mean summary statistic of firms by their respective price-setting strategy. Standard errors of the respective value are reported in parenthesis.

Table C.12: The Role of Financial Constraints - Interacting age and Size

	(1) $\Delta\text{CAPEX}_{t t+4}$ b/se
Young & Small \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	-0.27 (0.47)
Small & Old \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.05 (0.22)
Young & Medium/Large \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.46** (0.20)
Old & Medium/Large \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.35*** (0.11)
Young & Small = Young (p-value)	0.54
Young & Medium/Large = Small & Old (p-value)	0.15
Young & Medium/Large = Old & Medium/Large (p-value)	0.64
N	4,149

Note: This table reports the results from estimating model (5), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). We classify firms as young if they are less than 15 years old. We classify firms as small if they have fewer than 50 employees. Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. See Tables B.6 & B.7 for the respective median value of the firm characteristic and median sales growth surprises for each group.

Table C.13: The Role of Financial Constraints - Balance Sheet Measures

	(1)	(2)	(3)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se	b/se
Low Leverage \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.39** (0.18)		
Medium Leverage \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.44*** (0.15)		
High Leverage \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.18 (0.15)		
Low Liquidity \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.38** (0.17)	
Medium Liquidity \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.30** (0.14)	
High Liquidity \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.31** (0.15)	
Low Int. Burden \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$			0.51** (0.21)
Medium Int. Burden \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$			0.50*** (0.17)
High Int. Burden \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$			0.31 (0.19)
1st = 2nd (p-value)	0.82	0.74	0.97
2nd = 3rd (p-value)	0.26	0.94	0.39
N	3,833	3,783	2,876

Note: This table reports the results from estimating model (5), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively. See Tables B.6 & B.7 for the respective median value of the firm characteristic and median sales growth surprises for each group.

Table C.14: Heterogeneity by Margins

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
Lower Tercile (Margin) \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.40** (0.17)	0.04*** (0.01)
Middle Tercile (Margin) \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.37* (0.19)	0.03*** (0.01)
Upper Tercile (Margin) \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.69*** (0.26)	0.05*** (0.02)
Low = Medium (p-value)	0.89	0.41
Low = High (p-value)	0.35	0.45
N	2,471	2,133

Note: This table reports the results from estimating model (5), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.15: Heterogeneity by Markups

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
Lower Tercile (Markup) \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.08 (0.18)	0.04*** (0.01)
Middle Tercile (Markup) \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.38** (0.19)	0.03*** (0.01)
Upper Tercile (Markup) \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.49*** (0.18)	0.04*** (0.01)
Low = Medium (p-value)	0.26	0.57
Low = High (p-value)	0.11	0.73
N	3,017	2,588

Note: This table reports the results from estimating model (5), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.16: The Role of Sales Uncertainty

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{CAPEX}_{t t+4}$
	b/se	b/se
$\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$	0.31*** (0.08)	
Low Span \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.15 (0.15)
Medium Span \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.36** (0.15)
High Span \times $\text{FE}_{t-3}(\Delta\text{Sales}_{t-4 t})$		0.35*** (0.12)
Sales Span	0.02 (0.12)	
L4.Sales Span	-0.03 (0.11)	
L4.Sales Span		
1st = 2nd (p-value)		0.31
2nd = 3rd (p-value)		0.92
N	4,312	4,312

Note: This table reports the results from estimating models (2) (Columns (1) and (3)) as well (5) (Columns (2) and (4)), controlling for realised CAPEX growth, realised price growth, expected sales growth, and dummies for the respective interaction terms (omitted for brevity). Span refers to the difference in sales growth between the best and worst possible realisation. Standard errors clustered on firm as well as quarter \times industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.

Table C.17: Micro vs. Macro Surprises

	(1)	(2)
	$\Delta\text{CAPEX}_{t t+4}$	$\Delta\text{Price}_{t t+4}$
	b/se	b/se
Aggregate Component	2.56 (1.90)	1.21*** (0.17)
Idiosyncratic Component	5.88*** (1.59)	0.40*** (0.08)
$\mathbb{E}_{t+1}(\Delta\text{Sales}_{t t+4})$	1.01*** (0.11)	0.08*** (0.01)
$\Delta\text{CAPEX}_{t-4 t}$	-0.25*** (0.02)	-0.00 (0.00)
$\Delta\text{Price}_{t-4 t}$	-0.14 (0.37)	0.42*** (0.03)
Micro = Macro (p-value)	0.12	0.00
N	4,312	3,704

Note: This table reports the results from estimating model (2), decomposing the sales forecast error into a macro & industry specific component as well as a firm specific component by regressing the forecast errors on quarter x industry dummies. Standard errors clustered on firm as well as quarter x industry level. *, **, and *** indicate statistical significance at the 10%, 5%, and 1% level, respectively.